

Estimation and Control for Stochastic Regression Models

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Plan

- 1 Introduction
 - Goals
 - Weighted least squares algorithm
 - Adaptive tracking control
 - Optimization
- 2 Strong law of large numbers
- 3 Linear regression models with adaptive control
- 4 Almost sure central limit theorem
- 5 Functional regression models with adaptive control



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Consider the stochastic regression model

$$X_{n+1} = \theta^t \Phi_n + U_n + \varepsilon_{n+1}$$

- X_n → the system output,
- Φ_n → the regression vector,
- U_n → the adaptive control that can be chosen,
- ε_n → the dirven noise.

We have two goals

- Estimate the unknown parameter θ ,
- Control the dynamic of the process (X_n).



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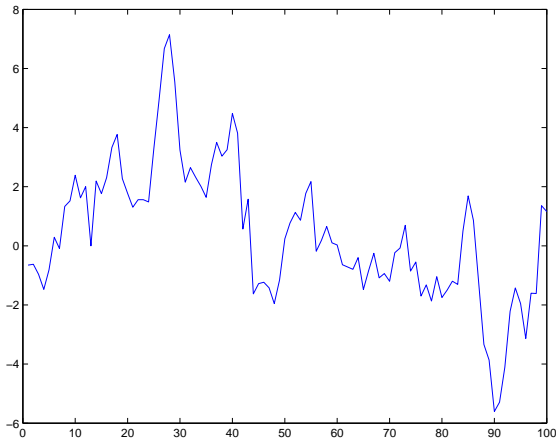
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Simulation of stable autoregressive process $|\theta| < 1$



$$X_{n+1} = \theta X_n + \varepsilon_{n+1}$$



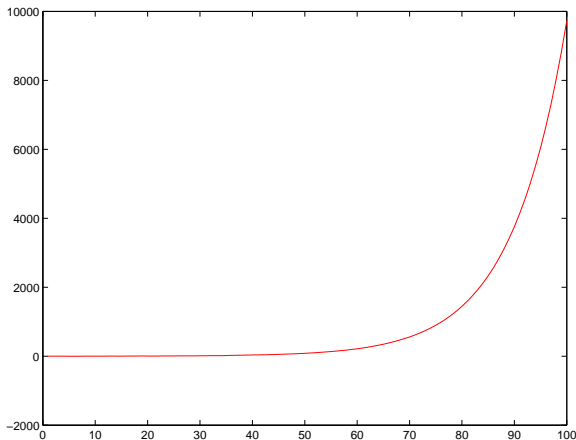
Introduction

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Almost sure central limit theorem
Functional regression models with adaptive control

Goals

Weighted least squares algorithm
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Simulation of explosive autoregressive process $|\theta| > 1$



$$X_{n+1} = \theta X_n + \varepsilon_{n+1}$$



The weighted least squares estimator $\hat{\theta}_n$ of θ minimises

$$\Delta_n(\theta) = \frac{1}{2} \sum_{k=0}^{n-1} a_k (X_{k+1} - U_k - \theta^t \Phi_k)^2.$$

Consequently,

$$\hat{\theta}_n = \mathbf{S}_{n-1}^{-1}(\mathbf{a}) \sum_{k=0}^{n-1} a_k \Phi_k (X_{k+1} - U_k),$$

$$\mathbf{S}_n(\mathbf{a}) = \sum_{k=0}^n a_k \Phi_k \Phi_k^t.$$

The **standard least squares estimator** is given by

$$a_n = 1.$$



Weighted least squares estimator

The **weighted least squares estimator** is given for $\gamma > 0$ by

$$\mathbf{a}_n = \left(\frac{1}{\log \mathbf{s}_n} \right)^{1+\gamma} \quad \text{where} \quad \mathbf{s}_n = \sum_{k=0}^n \|\Phi_k\|^2.$$

We always have the decomposition

$$\hat{\theta}_n - \theta = \mathbf{S}_{n-1}^{-1}(\mathbf{a}) \mathbf{M}_n(\mathbf{a})$$

$$\mathbf{M}_n(\mathbf{a}) = \sum_{k=0}^{n-1} \mathbf{a}_k \Phi_{k \in k+1}.$$



Adaptive tracking control

We wish to track, step by step, a given reference trajectory (x_n).
We make use of the **adaptive tracking control**

$$U_n = x_{n+1} - \hat{\theta}_n^t \Phi_n.$$

Hence, the closed-loop system is given by

$$x_{n+1} - x_{n+1} = \pi_n + \varepsilon_{n+1}$$

where

$$\pi_n = (\theta - \hat{\theta}_n)^t \Phi_n.$$



Optimization

We assume that (ε_n) satisfies the **law of large numbers**

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \varepsilon_k^2 = \sigma^2 \quad \text{a.s.}$$

where $\sigma^2 > 0$. We shall say that the **tracking is optimal** if

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n (X_k - x_k)^2 = \sigma^2 \quad \text{a.s.}$$



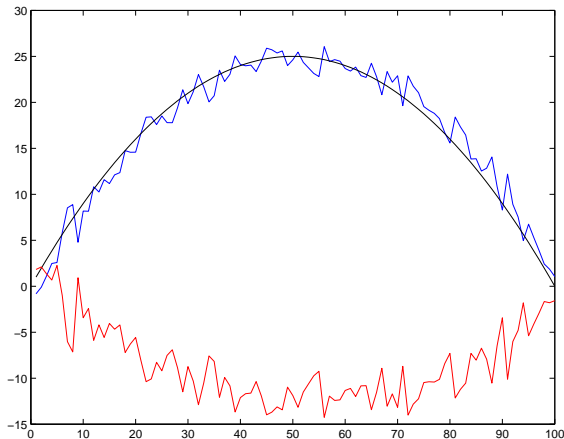
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$$X_{n+1} = \theta X_n + U_n + \varepsilon_{n+1}$$



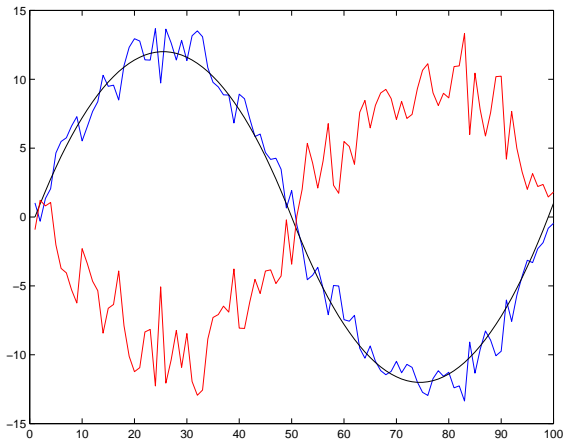
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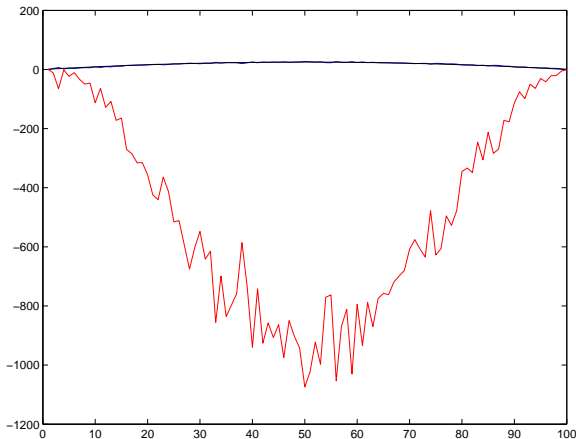
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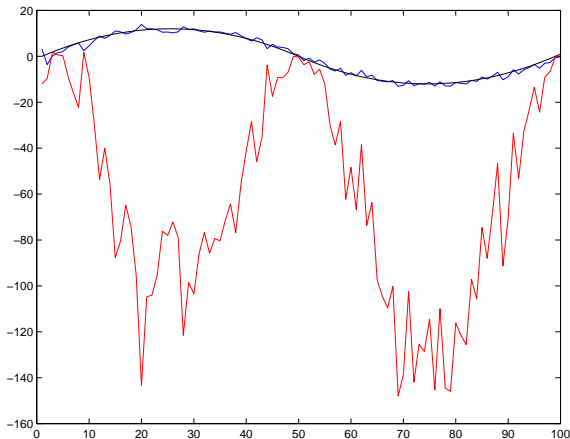
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Let (ε_n) be a sequence adapted to a filtration $\mathbb{F} = (\mathcal{F}_n)$ with

$$\mathbb{E}[\varepsilon_{n+1} | \mathcal{F}_n] = 0 \quad \text{and} \quad \mathbb{E}[\varepsilon_{n+1}^2 | \mathcal{F}_n] = \sigma^2 > 0.$$

For a scalar sequence (Φ_n) adapted to \mathbb{F} , we investigate the asymptotic behavior of the **martingale transform**

$$M_n = \sum_{k=1}^n \Phi_{k-1} \varepsilon_k.$$

The **explosion coefficient** associated to (Φ_n) is given by

$$f_n = \frac{\Phi_n^2}{S_n} \quad \text{where} \quad S_n = \sum_{k=0}^n \Phi_k^2.$$



First law of large numbers

In all the sequel, we assume that

$$\lim_{n \rightarrow \infty} s_n = +\infty \quad \text{a.s.}$$

Theorem (First LLN)

We have

$$(LLN) \quad \lim_{n \rightarrow \infty} \frac{M_n}{s_{n-1}} = 0 \quad \text{a.s.}$$

Remark. If (s_n) converges, then (M_n) also converges a.s.



Second law of large numbers

Theorem (Second LLN)

For $a > 2$, assume that

$$(A_1) \quad \sup_{n \geq 0} \mathbb{E} [|\varepsilon_{n+1}|^a | \mathcal{F}_n] < \infty \quad \text{a.s.}$$

Then, we have

$$\left(\frac{M_n^2}{S_{n-1}} \right) = \mathcal{O}(\log s_n) \quad \text{a.s.}$$

$$\sum_{k=1}^n f_k \left(\frac{M_k^2}{S_{k-1}} \right) = \mathcal{O}(\log s_n) \quad \text{a.s.}$$



Quadratic strong law

Theorem (Quadratic strong law)

If (A_1) holds and the explosion coefficient $f_n \rightarrow 0$ a.s., we have

$$(QSL) \quad \lim_{n \rightarrow \infty} \frac{1}{\log s_n} \sum_{k=1}^n f_k \left(\frac{M_k^2}{s_{k-1}} \right) = \sigma^2 \quad a.s.$$

Remark. The **QSL** is exactly the convergence of the moment of order 2 in the **ASCLT** for (M_n) .



Example

Let (ξ_n) be a sequence of **iid** random variables with $\mathbb{E}[\xi_n] = m$ and $\text{Var}(\xi_n^2) = \sigma^2$. If $S_n = \xi_1 + \xi_2 + \cdots + \xi_n$, we have

$$\text{(LLN)} \quad \lim_{n \rightarrow \infty} \frac{S_n}{n} = m \quad \text{a.s.}$$

$$\text{(QSL)} \quad \lim_{n \rightarrow \infty} \frac{1}{\log n} \sum_{k=1}^n \left(\frac{S_k - km}{k} \right)^2 = \sigma^2 \quad \text{a.s.}$$



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Consider the autoregressive process with adaptive control

$$\mathbf{X}_{n+1} = \theta^t \Phi_n + \mathbf{U}_n + \varepsilon_{n+1},$$

$$\Phi_n = (\mathbf{X}_n, \dots, \mathbf{X}_{n-p+1})^t.$$

We assume that the reference trajectory (x_n) satisfies

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n x_k^2 = \tau^2 \quad \text{a.s.}$$

where $\tau \geq 0$. For all $n \geq 0$, let

$$S_n = \sum_{k=0}^n \Phi_k \Phi_k^t.$$



Lemma (Bercu)

Assume that (A_1) holds. If $\ell = \sigma^2 + \tau^2$, we have

$$\lim_{n \rightarrow \infty} \frac{S_n}{n} = \ell I_p \quad \text{a.s.}$$

Theorem (Bercu)

If (A_1) holds, $\hat{\theta}_n$ **converges almost surely** to θ

$$\| \hat{\theta}_n - \theta \|^2 = \mathcal{O}\left(\frac{\log n}{n}\right) \quad \text{a.s.}$$

In addition, the **tracking is optimal**

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n (X_k - x_k)^2 = \sigma^2 \quad \text{a.s.}$$



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Theorem (Bercu)

If (A_1) holds, we have

$$(CLT) \quad \sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow{\mathcal{L}} \mathcal{N}\left(\mathbf{0}, \frac{\sigma^2}{\ell} I_p\right),$$

$$(LIL) \quad \limsup_{n \rightarrow \infty} \left(\frac{n}{2 \log \log n} \right) \|\hat{\theta}_n - \theta\|^2 = \frac{\sigma^2}{\ell} \quad \text{a.s.}$$

$$(QSL) \quad \lim_{n \rightarrow \infty} \frac{1}{\log n} \sum_{k=1}^n \|\hat{\theta}_k - \theta\|^2 = \frac{\sigma^2}{\ell} \quad \text{a.s.}$$



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Central limit theorem

Let (ξ_n) be a sequence of **iid** random variables with $\mathbb{E}[\xi_n] = m$ and $\text{Var}(\xi_n) = \sigma^2$. If $S_n = \xi_1 + \xi_2 + \dots + \xi_n$, we have

$$(CLT) \quad \frac{S_n - nm}{\sqrt{n}} \xrightarrow{\mathcal{L}} \mathcal{N}(0, \sigma^2).$$

In other words, for any function **h bounded continuous**,

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[h \left(\frac{S_n - nm}{\sqrt{n}} \right) \right] = \int_{\mathbb{R}} h(x) dG(x)$$

where G stands for the Gaussian measure $\mathcal{N}(0, \sigma^2)$.



Almost sure central limit theorem

We also have

$$(ASCLT) \quad \frac{1}{\log n} \sum_{k=1}^n \frac{1}{k} \delta \left(\frac{S_k - km}{\sqrt{k}} \right) \implies \mathbf{G} \quad \text{a.s.}$$

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We have already seen the **LLN** for the **martingale transform**

$$M_n = \sum_{k=1}^n \Phi_{k-1} \varepsilon_k.$$

The **explosion coefficient** associated with (Φ_n) is given by

$$f_n = \frac{\Phi_n^2}{S_n} \quad \text{avec} \quad S_n = \sum_{k=0}^n \Phi_k^2.$$



Theorem (Brown, Chaabane, Lifshits)

If (A_1) holds and the explosion coefficient $f_n \rightarrow 0$ a.s., we have

$$(CLT) \quad \frac{M_n}{\sqrt{s_{n-1}}} \xrightarrow{\mathcal{L}} \mathcal{N}(0, \sigma^2).$$

In addition, if

$$\sum_{n=1}^{\infty} f_n^\gamma < \infty \quad \text{a.s.}$$

for some $\gamma > 0$, then we also have

$$(ASCLT) \quad \frac{1}{\log s_n} \sum_{k=1}^n f_k \delta \left(\frac{M_k}{\sqrt{s_{k-1}}} \right) \Longrightarrow \mathbf{G} \quad \text{a.s.}$$



Powers of martingales

For any function **h bounded continuous**, we have

$$\lim_{n \rightarrow \infty} \frac{1}{\log s_n} \sum_{k=1}^n f_k h\left(\frac{M_k}{\sqrt{s_{k-1}}}\right) = \int_{\mathbb{R}} h(x) dG(x) \quad \text{a.s.}$$

Definition. We shall say that (M_n) satisfies a **PASCLT** if this convergence holds for all **polynomial function h** .

Goal. Establish a **PASCLT** in order to study the stability of controlled functional regression models.



Let

$$v_n(p) = \frac{s_n^p - s_{n-1}^p}{s_n^p}.$$

Theorem (Bercu)

For some $p \geq 1$ and $a > 2p$, assume that

$$(A_p) \quad \sup_{n \geq 0} \mathbb{E} [|\varepsilon_{n+1}|^a | \mathcal{F}_n] < \infty \quad \text{a.s.}$$

Then, we have

$$\left(\frac{M_n^2}{s_{n-1}} \right)^p = \mathcal{O}(\log s_n) \quad \text{a.s.}$$

$$\sum_{k=1}^n v_k(p) \left(\frac{M_k^2}{s_{k-1}} \right)^p = \mathcal{O}(\log s_n) \quad \text{a.s.}$$



Theorem (Bercu)

If (A_p) holds and the explosion coefficient $f_n \rightarrow 0$ a.s., we have

$$\lim_{n \rightarrow \infty} \frac{1}{\log s_n} \sum_{k=1}^n f_k \left(\frac{M_k^2}{s_{k-1}} \right)^p = \frac{\sigma^{2p} (2p)!}{2^p p!} \quad \text{a.s.}$$

Theorem (Bercu-Fort)

Assume that (A_p) holds for all $p \geq 1$ and $f_n \rightarrow 0$ a.s. Then, (M_n) satisfies the **PASCLT**

$$\frac{1}{\log s_n} \sum_{k=1}^n f_k \delta \left(\frac{M_k}{\sqrt{s_{k-1}}} \right) \implies \mathbf{G} \quad \text{a.s.}$$



Explosive martingales

For all $p \geq 1$, let

$$\sigma_n(p) = \mathbb{E}[\varepsilon_{n+1}^p | \mathcal{F}_n] \quad \text{a.s.}$$

Theorem (Bercu)

Assume that (A_p) holds and, for all $2 \leq q \leq 2p$, $\sigma_n(q) \rightarrow \sigma(q)$ a.s. where $\sigma(q) = 0$ if q is odd. Also assume that $\mathbf{f}_n \rightarrow \mathbf{f}$ a.s. where $0 < \mathbf{f} < 1$. Then, we have

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \left(\frac{M_k^2}{\mathbf{s}_{k-1}} \right)^p = I(p, \mathbf{f}) \quad \text{a.s.}$$



Gaussian limit distribution

The limit $l(p, \mathbf{f})$ is given by

$$l(p, \mathbf{f}) = \frac{1}{1 - (1 - \mathbf{f})^p} \sum_{k=1}^p C_{2p}^{2k} \mathbf{f}^k (1 - \mathbf{f})^{p-k} \sigma(2k) l(p-k, \mathbf{f}).$$

This expression does not depend on \mathbf{f} iff, for all $1 \leq k \leq p$,

$$\sigma(2k) = \frac{\sigma^{2k} (2k)!}{2^k k!}.$$

In that particular case, we have

$$l(p, \mathbf{f}) = \frac{\sigma^{2p} (2p)!}{2^p p!} = l(p).$$



Explosive martingales

Theorem (Bercu-Fort)

Assume that (A_p) holds for all $p \geq 1$ and $\mathbf{f}_n \rightarrow \mathbf{f}$ a.s. where $0 < \mathbf{f} < 1$. For all $p \geq 1$, if $l(\mathbf{p}, \mathbf{f}) = l(\mathbf{p})$, then (M_n) satisfies the **PASCLT**

$$\frac{1}{n} \sum_{k=1}^n \delta \left(\frac{M_k}{\sqrt{\mathbf{S}_{k-1}}} \right) \implies \mathbf{G} \quad \text{a.s.}$$



Stable autoregressive process $|\theta| < 1$

Consider the stable autoregressive process

$$X_{n+1} = \theta X_n + \varepsilon_{n+1}.$$

If (A_1) holds, we have $f_n \rightarrow 0$,

$$\lim_{n \rightarrow \infty} \frac{S_n}{n} = \frac{\sigma^2}{(1 - \theta^2)} \quad \text{a.s.}$$

In addition, $\hat{\theta}_n \rightarrow \theta$ a.s. and

$$\sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow{\mathcal{L}} \mathcal{N}(0, 1 - \theta^2).$$



If (A_p) holds for all $p \geq 1$, we have the **PASCLT**

$$\frac{1}{\log n} \sum_{k=1}^n \frac{1}{k} \delta_{\sqrt{k}(\hat{\theta}_k - \theta)} \implies \mathcal{N}(0, 1 - \theta^2) \quad \text{a.s.}$$

In particular, for all $p \geq 1$, we have

$$\lim_{n \rightarrow \infty} \frac{1}{\log n} \sum_{k=1}^n k^{p-1} (\hat{\theta}_k - \theta)^{2p} = \frac{(1 - \theta^2)^p (2p)!}{2^p p!} \quad \text{a.s.}$$



Explosive autoregressive process $|\theta| > 1$

If (H_1) holds, $\theta^{-n}X_n$ converges a.s. to the random variable

$$Y = X_0 + \sum_{k=1}^{\infty} \theta^{-k} \varepsilon_k.$$

In addition, $f_n \rightarrow (\theta^2 - 1)/\theta^2$,

$$\lim_{n \rightarrow \infty} \frac{S_n}{\theta^{2n}} = \frac{\theta^2 Y^2}{(\theta^2 - 1)} \quad \text{a.s.}$$

Consequently, $\hat{\theta}_n \rightarrow \theta$ a.s. Moreover, if (ε_n) is **gaussian** and \mathcal{C} stands for the Cauchy distribution

$$|\theta|^n (\hat{\theta}_n - \theta) \xrightarrow{\mathcal{L}} \mathcal{C}.$$



If (ε_n) is **gaussian**, we have the **PASCLT**

$$\frac{1}{n} \sum_{k=1}^n \delta_{|\theta|^k(\hat{\theta}_k - \theta)} \implies \mathcal{N}\left(\mathbf{0}, \frac{\sigma^2(\theta^2 - 1)}{\gamma^2}\right) \quad \text{a.s.}$$

In particular, for all $p \geq 1$, we have

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n (|\theta|^k(\hat{\theta}_k - \theta))^{2p} = \frac{\sigma^{2p}(\theta^2 - 1)^p (2p)!}{\gamma^{2p} 2^p p!} \quad \text{a.s.}$$



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Consider the functional autoregressive model of order $d \geq 1$

$$X_{n+1} = \theta f(X_n, \dots, X_{n-d+1}) + U_n + \varepsilon_{n+1}.$$

We estimate θ by the **standard least squares estimator**

$$\hat{\theta}_n - \theta = \frac{M_n}{S_{n-1}} \quad \text{with} \quad M_n = \sum_{k=1}^n \Phi_{k-1} \varepsilon_k.$$

We choose the **adaptive tracking control**

$$U_n = x_{n+1} - \hat{\theta}_n \Phi_n$$

where $\Phi_n = f(X_n, \dots, X_{n-d+1})$.



The functional class $\mathcal{C}(a, b)$

Let $\mathcal{C}(a, b)$ with $a, b \in \mathbb{N}$ and $a \geq 1$ be the class of functions f from \mathbb{R}^d to \mathbb{R} such that, for all $x \in \mathbb{R}^d$,

$$c_1 + c_2 \|x\|^b \leq |f(x)| \leq c_3 + c_4 \|x\|^a$$

where $b \geq 1$ if $c_1 = 0$ and $b \geq 0$ otherwise.



Corollary (Bercu-Portier)

Assume that (H_a) holds and $f \in \mathcal{C}(a, b)$ with $a < 4$. Then, we have $\hat{\theta}_n \rightarrow \theta$ a.s. and

$$(\hat{\theta}_n - \theta)^2 = \mathcal{O}\left(\frac{\log n}{n}\right) \quad \text{a.s.}$$

For all $1 \leq p \leq a$, the **tracking is stable of order p**

$$\limsup_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n (X_k - x_k)^{2p} < \infty \quad \text{a.s.}$$

If $\sigma_n(2p) \rightarrow \sigma(2p)$ a.s., the **tracking is optimal of order p**

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n (X_k - x_k)^{2p} = \sigma(2p) \quad \text{a.s.}$$

The natural hypothesis (H_a)

Denote by $\mathcal{P}(a)$ the polynomial algebra with d variables and **total degree** $\leq a$ with $a \geq 1$. We assume that $f^2 \in \mathcal{P}(2a)$ together with

$$(H_a) \quad \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=d}^n f^2(\varepsilon_k + x_k, \dots, \varepsilon_{k-d+1} + x_{k-d+1}) = \ell \quad \text{a.s.}$$

where $\ell > 0$. Under (A_a) and (H_a) with $a < 4$, we can prove

$$\lim_{n \rightarrow \infty} \frac{S_n}{n} = \ell \quad \text{a.s.}$$



Corollary (Bercu-Portier)

Under (A_a) and (H_a) with $a < 4$, we have $\hat{\theta}_n \rightarrow \theta$ a.s. and

$$(CLT) \quad \sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow{\mathcal{L}} \mathcal{N}\left(\mathbf{0}, \frac{\sigma^2}{\ell}\right),$$

$$(LIL) \quad \limsup_{n \rightarrow \infty} \left(\frac{n}{2 \log \log n} \right) (\hat{\theta}_n - \theta)^2 = \frac{\sigma^2}{\ell} \quad \text{a.s.}$$

Moreover, for all $1 \leq p \leq a$, we also have

$$\lim_{n \rightarrow \infty} \frac{1}{\log n} \sum_{k=1}^n k^{p-1} (\hat{\theta}_k - \theta)^{2p} = \frac{\sigma^{2p} (2p)!}{\ell^p 2^p p!} \quad \text{a.s.}$$



Polynomial autoregressive processes of order 2

Assume that $\mathbf{x}_n \rightarrow \mathbf{0}$ and $\sigma_n(\mathbf{p}) \rightarrow \sigma(\mathbf{p})$ a.s. for all $1 \leq p \leq 4$. Consider the polynomial autoregressive processes

$$(1) \quad X_{n+1} = \theta X_n^2 + U_n + \varepsilon_{n+1},$$

$$(2) \quad X_{n+1} = \theta X_n(1 - X_n) + U_n + \varepsilon_{n+1},$$

$$(3) \quad X_{n+1} = \theta X_n X_{n-1} + U_n + \varepsilon_{n+1}.$$

Then, the corollary holds with $\ell(1) = \sigma(4)$,

$$\ell(2) = \sigma(4) - 2\sigma(3) + \sigma(2), \quad \ell(3) = \sigma(2)^2.$$



Polynomial autoregressive processes of order 3

Assume that $\mathbf{x}_n \rightarrow \mathbf{0}$ and $\sigma_n(\mathbf{p}) \rightarrow \sigma(\mathbf{p})$ a.s. for all $1 \leq p \leq 6$. Consider the polynomial autoregressive processes

$$(4) \quad X_{n+1} = \theta X_n^3 + U_n + \varepsilon_{n+1},$$

$$(5) \quad X_{n+1} = \theta X_n^2(1 - X_n) + U_n + \varepsilon_{n+1},$$

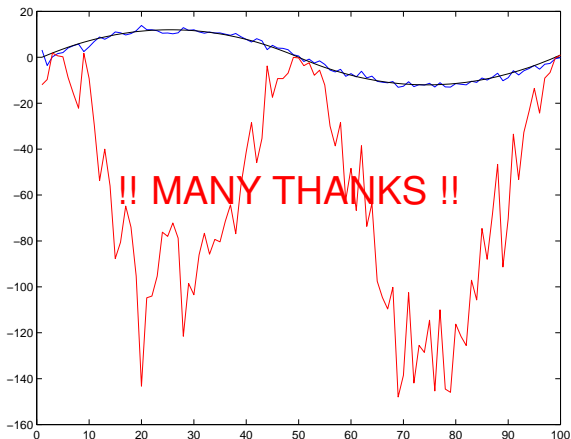
$$(6) \quad X_{n+1} = \theta X_n^2 X_{n-1} + U_n + \varepsilon_{n+1}.$$

Then, the corollary holds with $\ell(4) = \sigma(6)$,

$$\ell(5) = \sigma(6) - 2\sigma(5) + \sigma(4), \quad \ell(6) = \sigma(4)\sigma(2).$$



Simulation of controlled autoregressive process



$$X_{n+1} = \theta X_n^2 + U_n + \varepsilon_{n+1}$$

