"Shape Invariant" curves estimation and application to prediction problems

J-M. LOUBES

Institut de Mathématiques de Toulouse

IHP Rencontres INRIA-AMIES-EDF 2012

Outline

Shape invariant model

Curve Registration

Prediction using shape invariant model

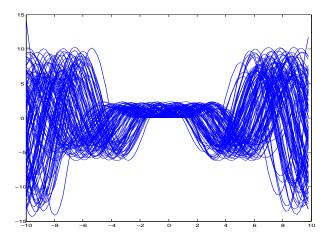
A practical Issue in functional analysis

- Functional data usually convey a general information, which reflects the inner structure of the observations
- but also different sources of variation which blurs the data and prevent the use of the Euclidean mean.

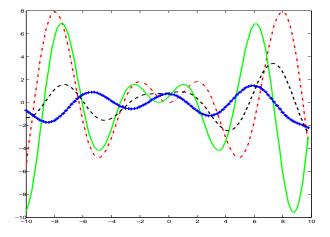
Common in biology, medicine (mixed effect model), economics or sociology (panel data) and also in electric charge.

Appears when individuals may differ slightly from a pattern which represents the shape of the observations and break the Euclidean structure of the data.

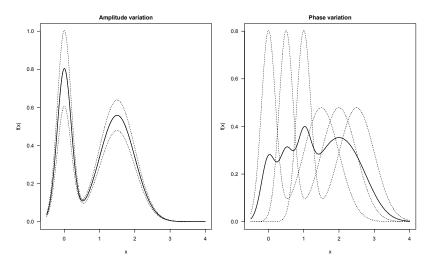
Simulations: Do we trust the mean?



Simulations: Do we trust the mean?

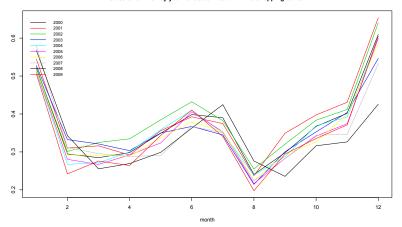


Example: in oligonucleotide array

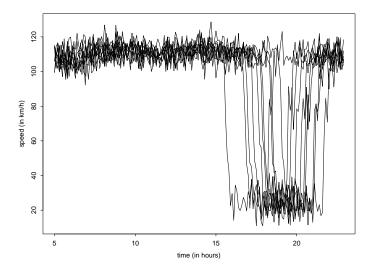


Example: income of stores from a mall

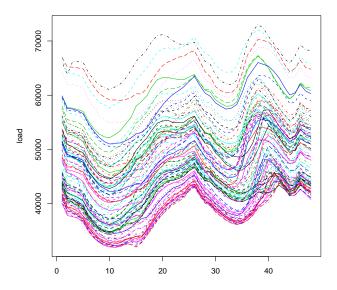
Sales/size index by year for sector 2000 in El123 shopping center



Example: traffic jam on a motorway



Example: EDF daily load



Objectives of the statistical studied

Main issues:

- Recover the shape or pattern of the data from the warped observations
- Estimate the individuals deformations
- Understand the deformations: cluster the data according to the distance given by the observations, i.e the geometry of the data.
- Forecast the future deformations using auxiliary information
- ⇒ different framework : shape analysis of curves viewed as a manifold with an inner geometry.

General Model

The regression model:

$$Y_{i,j} = f_i^*(t_{ij}) + \sigma \epsilon_{i,j}, \quad i = 1,...,n, \ j = 1,...,J.$$

where

- f_i^* models the j^{th} signal (unknown);
- t_{ij} the observation points (known).
- $\epsilon_{i,j}$ is white noise (unknown), and σ variance (unknown)

Assumption: There exists a **common shape** of the signal f^* and warping operators Φ_j ,

$$f_i^{\star} = \Phi_j f^{\star}, \quad j = 1, \ldots, J.$$

Aim: Estimation of the deformations and the template f^*

Methodology

inverse problem of regression for unknown operator

=> Need for a Model for the warping operator

Parametric Modeling and Semiparametric Statistics

The warping operator can be **parametrized** by $\theta \in \Theta \subset \mathbb{R}^d$.

$$orall j = 1, \dots, J, \quad \Phi_j = \Phi_{ heta_j}$$
 $Y_{ij} = \Phi_{ heta_j^*}[f^*](t_{ij}) + \sigma \epsilon_{ij}$

- f^* is the main feature
- $\Phi_{\theta^{\star}}[.]$ is a parametric warping operator
- θ_i^{\star} local warping parameters

Objective : estimate f^* by estimating the θ_i^* 's

Semiparametric framework with unknown distribution $\mathbb{P}_{f^{\star},\theta^{\star}}$

A class of deformations

$$\theta = (a \quad b \quad v)'$$

$$\Phi_{\theta}: f(\cdot) \to af(\cdot - b) + v$$

Parametric Modeling and Semiparametric Statistics

The warping operator can be **parametrized** by $\theta \in \Theta \subset \mathbb{R}^d$.

$$orall j = 1, \dots, J, \quad \Phi_j = \Phi_{ heta_j}$$
 $Y_{ij} = \Phi_{ heta_j^*}[f^*](t_{ij}) + \sigma \epsilon_{ij}$

- f* is the main feature
- $\Phi_{\theta^*}[.]$ is a parametric warping operator
- θ_i^{\star} local warping parameters

Objective : estimate f^* by estimating the θ_i^* 's

Semiparametric framework with unknown distribution $\mathbb{P}_{f^{\star},\theta^{\star}}$

A class of deformations

$$\theta = (a \quad b \quad \upsilon)'$$

$$\Phi_{\theta}: f(\cdot) \to af(\cdot - b) + v$$

Parametric Modeling and Semiparametric Statistics

The warping operator can be **parametrized** by $\theta \in \Theta \subset \mathbb{R}^d$.

$$orall j = 1, \dots, J, \quad \Phi_j = \Phi_{ heta_j}$$
 $Y_{ij} = \Phi_{ heta_j^*}[f^*](t_{ij}) + \sigma \epsilon_{ij}$

- f* is the main feature
- $\Phi_{\theta^*}[.]$ is a parametric warping operator
- θ_i^* local warping parameters

Objective : estimate f^* by estimating the θ_i^* 's

Semiparametric framework with unknown distribution $\mathbb{P}_{f^{\star},\theta^{\star}}$

A class of deformations

$$\theta = (a \quad b \quad v)'$$

$$\Phi_{\theta}: f(\cdot) \to af(\cdot - b) + v$$

Curve Alignment

$$Y_{i,j} = a_j^* f^*(t_{ij} - b_j^*) + v_j^* + \sigma_j^* \epsilon_{i,j}, \quad j = 1, \ldots, J, \ i \in \{1, \ldots, n\}^d.$$

Estimation of the parameters $\mathbf{b}_{i}^{\star} \in \mathbb{R}^{d}, \ \mathbf{a}_{i}^{\star} \in \mathbb{R}^{*}, \ v_{i}^{\star} \in \mathbb{R}, \ \mathbf{j} = 1, \ldots, \mathbf{J}$

Lawton, W.M., Sylvestre, E.A. et Maggio, M.G. (1972): iterative method

Kneip, A. et Gasser, T.(1988): consistency of the SEMOR method.

(**SEMOR**) based on the polynomial approximation of f^* .

Gamboa, F., Loubes, J-M. and Maza, E. (EJS 2007): Semi-parametric estimation of shifts.

Bigot J., Loubes, J-M. and Vimond, M. (PTRF 2012): Semiparametric estimation of rigid transformations on compact Lie groups,

Mathematical model

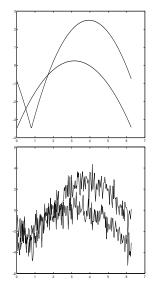
$$Y_{ij} = a_j^* f^*(t_i - b_j^*) + v_j^* + \sigma_j^* \epsilon_{ij}, i \in I_n = \{1..n\}^d,$$

Assumptions:

- $f^* \ 2\pi \mathbb{Z}^d$ -periodic continuous
- $t_i = 2\pi(i)/n \in [0, 2\pi[d],$
- $(\epsilon_{i,j})_{ij}$ are i.i.d, $\mathbb{E}\epsilon_{i,j}=0$, $\mathbb{E}\epsilon_{i,j}^2=1$,
- the variance of noise $\sigma_i^{*2} = \sigma^{*2}$,

Method:

- **1** Estimate the parameters θ^*
- Invert the estimated operator to estimate the shape f*



Identifiability Constraint

Two sets of identifiability constraints are considered:

 a natural parametrization: we consider one of the signal as a reference,

$$b_1^* = 0, \quad a_1^* = 1 \quad \text{and} \quad v_1^* = 0$$

2 an alternative parametrization:

$$b_1^* = 0, \quad \sum_{j=1}^J a_j^{*2} = J \quad \text{and} \quad a_1 > 0, \quad (1)$$
 $c_0(f^*) = \int_{\mathcal{D}} f(t) \frac{dt}{(2\pi)^d} = 0 \quad (2)$

- (2) leads us to an asymptotically independent estimators,
- (1) leads us to an asymptotically efficient estimators (profile likelihood) in the gaussian case,

Construction of a registration method

A new criterion : alignment of an individual warped curve onto the mean of all the warped curves $\Phi = \Phi_{\theta}$: parametric model for deformations => Semiparametric statistics

$$\theta = (a, b, v)', \quad \Phi_{\theta} : f(\cdot) \to af(\cdot - b) + v$$

• For a **candidate** θ , compute the deformations

$$g_j(\theta,x) = \Phi_{\theta}^{-1} \circ f_j^{\star}(x) = \Phi_{\theta_j}^{-1} \circ f_j^{\star}(x)$$

Registration Criterion

$$M_0(\theta)
ightarrow rac{1}{J} \sum_{j=1}^J \left\| g_j(\theta,x) - rac{1}{J} \sum_{j'=1}^J g_{j'}(\theta,x)
ight\|_{L^2}^2,$$

M-estimators of the parameters $\hat{\theta}$ using en empirical version $M_n(\theta)$ of $M_0(\theta)$

Semi-parametric framework for translations

Assumption : $\Phi_{\theta}[f](t) = f(t - \theta)$

$$Y_{ij} = f^*(t_{ij} - \theta_j^*) + \sigma \epsilon_{ij}$$

Fourier Transform (DWT) => Equivalent Observation Model

$$d_{jl} = \exp(-il\alpha_{j}^{\star})c_{l}(f) + w_{jl},$$

$$I = -\frac{n-1}{2}, \dots, \frac{n-1}{2}, j = 1, \dots, J.$$

 $w_{jl}=w_{jl}^{\times}+iw_{jl}^{Y}, \quad w_{jl}^{\{\times,y\}}\sim\mathcal{N}(0,\frac{\sigma^2}{n}), \text{ i.i.d.}$: observation noise $c_l(f),\ l\in\mathbb{Z}$: Fourier coefficients of f (unknown). $\alpha_j^{\star}=\frac{2\pi}{T}\theta_j^{\star}$ and $\alpha^{\star}=(\alpha_j^{\star})_{j=1,\dots,J_n}^{\prime}$ are warping parameters to estimate using d_{il} .

Construction of Contrast

Idea : Aligning one shifted coefficient to mean of the others

$$\tilde{d}_{jl}(\alpha) = \exp(il\alpha_j)d_{jl} = \exp(il[\alpha_j - \alpha_j^{\star}])c_l(f) + \exp(il\alpha_j)w_{jl}$$

$$\forall \alpha \in \mathbb{R}^{J_n}, \ \underline{M_n(\alpha)} = \sum_{l=-\frac{n-1}{2}}^{\frac{n-1}{2}} \frac{1}{J_n} \sum_{j=1}^{J_n} |\tilde{d}_{jl} - \frac{1}{J_n} \sum_{j=1}^{J_n} \tilde{d}_{jl}|^2.$$

• Smoothing Sequence δ_I , such that $\sum_l \delta_l^2 c_l^2 < +\infty$

$$\forall \alpha \in \mathbb{R}^{J_n}, \ \underline{M_n(\alpha)} = \sum_{l=-\frac{n-1}{2}}^{\frac{n-1}{2}} \frac{1}{J_n} \sum_{j=1}^{J_n} \delta_l^2 |\tilde{d}_{jl} - \frac{1}{J_n} \sum_{j=1}^{J_n} \tilde{d}_{jl}|^2.$$

M-estimation estimator

$$\hat{\alpha}_n = \arg\min_{\alpha \in [0,T]^{J_n} \cap \mathcal{A}_1} M_n(\alpha)$$

Convergence of the estimator

Assumptions

- (I) Convergence : $(\delta_I c_I(f))_I \in \ell^2(\mathbb{Z})$ and $\sum_{l \neq 0} \delta_I^2 c_I^2(f) \neq 0$ very weak
- (II) Asymptotic Normality : $\sum (\delta_I I)^2 c_I^2(f) < +\infty$, $\sum (\delta_I I)^4 c_I^2(f) < +\infty$ et $\sum_{|I| < n/2} I^4 \delta_I^4 = o(n^2)$

Theorem (Gamboa-Loubes-Maza (2007))

Under Assumptions (I), $\hat{\alpha}_n$ converges in Probability to α^* .

Proof using **Standard Technics** in M-estimation theory:

- $M_n(\alpha) \xrightarrow{\mathbf{P}_{\alpha^{\star}}} K(\alpha, \alpha^{\star})$ with minimum $\alpha = \alpha^{\star}$
- \mathcal{A} is compact, $\alpha \to M_n(\alpha)$ and $\alpha \to K(\alpha, \alpha^*)$ continuous.
- $\lim_{n} \mathbf{P}_{\alpha^{\star}} [\sup_{|\alpha \beta| \le \eta_{k}} |M_{n}(\alpha) M_{n}(\beta)| > \epsilon_{k}] = 0.$

Asymptotic Normality

Theorem

Under Assumptions (II), the estimate $\hat{\alpha}_n$ is such that

$$\sqrt{n}(\hat{\alpha}_n - \alpha^*) \xrightarrow{\mathcal{L}} \mathcal{N}(0, \Gamma),$$

with
$$\Gamma = \frac{(J_n-1)\sum_{l\in\mathbb{Z}} l^2 \delta_l^2}{J_n^6(J_n-2)^2(\sum_{l\in\mathbb{Z}} \delta_l l^2 |c_l(f)|)^2} V(J).$$

For
$$\delta_l \approx c_l(f)$$
, if $\sum_l l^2 c_l(f)^2 < +\infty$

$$\sqrt{n\sum_{l}l^{2}c_{l}(f)^{2}}(\hat{\alpha}_{n}-\alpha^{\star})\stackrel{\mathcal{L}}{\longrightarrow}\mathcal{N}\left(0,V(J)\right)$$

 \rightarrow optimality of Warping Procedure with $\hat{\alpha}_n$.

Complete Model $\Phi_{\theta}: f(\cdot) \to af(\cdot - b) + v$

$$d_{j0} = a_j^* c_0(f^*) + \upsilon_j^* + w_{j0}$$

$$d_{jl} = a_j^* \exp(-il\frac{2\pi}{T}b_j^*)c_l(f^*) + w_{jl}, \quad l \neq 0.$$

Inverting the operator for a candidate $\theta = (\alpha \quad b \quad v)'$

$$\widetilde{d}_{jl}(\theta) = rac{\exp(ilrac{2\pi}{T}b_j)}{a_j}d_{jl}, \ l \neq 0, \quad \widetilde{d}_{j0}(\theta) = rac{1}{a_j}(d_{j0} - v_j).$$

$$\widehat{\theta}_n = \arg\min_{\theta} M_n(\theta) \longrightarrow \theta^*$$

Shape invariant curve estimator of f^*

$$\widehat{f}(t_i) = rac{1}{J} \sum_{i=1}^J \Phi_{\widehat{ heta}_j}^{-1} [Y_{ij}]$$

Other kind of deformations

Let v a vector field, $v_j, j=1,\ldots,J$ ϕ_j is a diffeomorphism defined as a solution of PDE

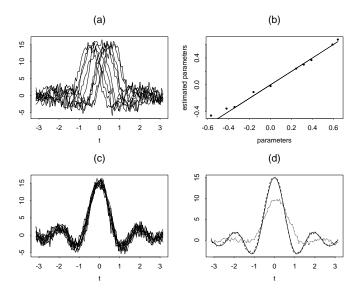
$$\frac{\partial}{\partial t}\phi_j = \mathsf{v}_j \circ \phi$$

Model proposed by Trouvé and Youness for image transformations and used for curve warping

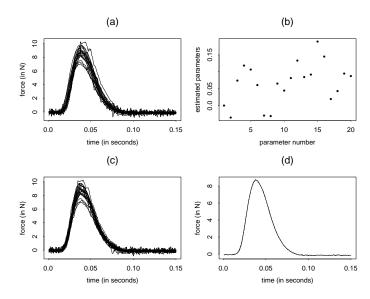
- Drawback: Fourier transform has not the same nice form but direct minimization of the quadratic criterion is still feasible;
- Advantage: allows non linear warping families

Bigot, Gadat, Loubes (2011)

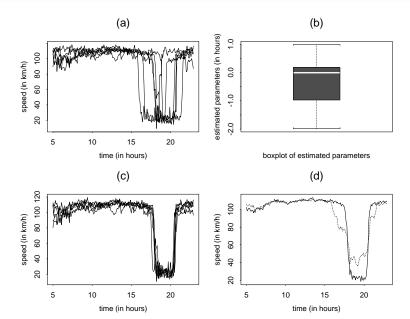
Simulations with sinusoidal functions



Pinch Force Data



Velocities of Cars

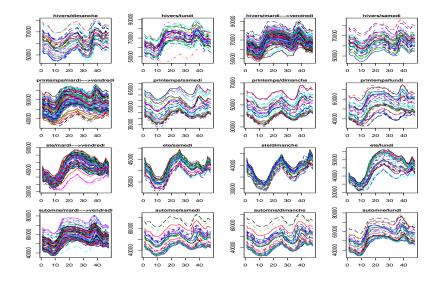


Velocities of Cars

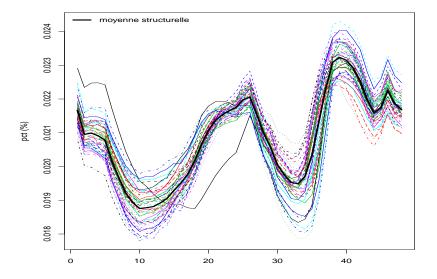


It works: good models of road trafficking behaviours joint work with $\mbox{{\bf V-Trafic}}$ i-phone application with Meteo as auxiliary variable for prediction

Clustering Load curves with warping distance



Registered Load curves



Conclusions: (partial)

- The shape represents a mean behaviour of a large number of individuals sharing the same behaviour with some fluctuations.
- Provides a different distance based on deformations
 - To aggregate the individual curves and provide a structural mean curve
 - To cluster the individuals with respect to the pseudo-distance to the pattern
 - To analyze the deformations with geometric PCA
- Based on sharp estimation of the deformations which relies on semi-parametric technics
- If other variables are available ... use them! And obtain better results and enable to **predict** using auxiliary variables

Incorporate additional information $X \in \mathcal{X}$

Framework: the outcome depends on **exogenous information** which explain the individual deviations from the pattern.

$$Y_{ij} = \Phi_{\theta_i^{\star}}[f^{\star}](t_i) + \sigma \varepsilon_{ij} = F(X_j, t_i).$$

 X_j are observed parameters which characterize the behaviour of the individual j

• Objective : model the relationship and estimate it

$$\theta$$
: $X_i \longrightarrow \theta(X_i) = \theta_i^*$

Forecast with shape invariant model

$$\widehat{f}_j = \Phi_{\widehat{\theta}(X_i)}[\widehat{f}]$$

Example: confidence index for income analysis, characteristics of the observation day (holidays, week end, standard) for vehicle speed, type of days or temperature evolution for electric load.

Building a surface response

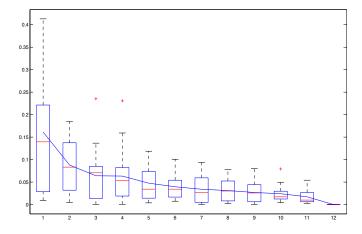
- lacksquare Data : $X_j,\,j=1,\ldots,J$ and the outcome Y_{ij}
- ② Estimation of the parameters $\widehat{ heta}_j \in \mathbb{R}^3$ and \widehat{f} of the shape f^\star
- ③ $\widehat{\theta}(X_j)$ is viewed as a sample of a random process $\widehat{\theta}(.)$ observed at random locations. $X_j, \quad j=1,\ldots,J$
 - \Rightarrow Prediction using standard **Kriging method** choosing a proper covariance structure K.

$$x_o \notin \{X_1, \dots, X_j\} \quad x_o \mapsto \tilde{\theta}(x_0)$$

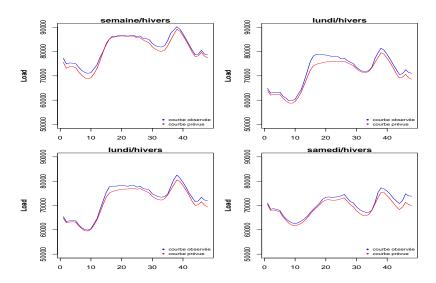
• Forecast : the predicted shape built by the deformations induced by the parameters x_0 .

$$\widehat{f}_0 = \Phi_{\widetilde{\theta}(\mathbf{x}_0)}[\widehat{f}]$$

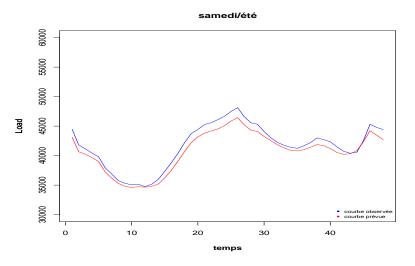
Forecasting Sales



Forecasting Load curves J+1



Forecasting Load curves J + 1



Remarks on the forecast

- Semiparametric framework enables efficient and sharp estimation of the underlying $\widehat{\theta}(.)$.

 Necessary since the prediction are based on a good approximation result.
- Choice of the Kernel (covariance) induces a distance on the covariates X
 For curves (temperature), the kernel must handle functional distance and may depend on a tuning parameter (see work by Sapatinas 2012).
- Sensitivity Analysis to analyze the influence of the parameters X.
- The deformations should be local ... for the moment work on parts of curves but work still in progress
- If scale issues: transform the problem into forecast of the daily distribution of consumption (deformation of densities) and the daily consumption (regression)